

A Probabilistic Model for Photon Pileup in Imaging Detectors

17th IACHEC Meeting
Osaka, Japan

Robert Zimmerman¹
with David A. van Dyk¹ and Vinay Kashyap²

¹Imperial College London

²Harvard & Smithsonian

May 13, 2025

Pileup

- **Pileup** occurs when multiple photons strike the same detector region during a single frame
- The detector cannot resolve them as separate events
- Their combined charge is interpreted as a *single* event:
 - ▶ Total energy is roughly the sum of the individual photon energies
 - ▶ Resulting charge cloud may be irregular → *bad grade*
- These effects distort the observed spectrum and reduce usable counts



Figure: Pileup distorts the observed image, producing a “hole” in bright regions where events are lost due to bad grades

Why Model Pileup Probabilistically?

- In high-energy regimes, photon coincidence distorts observed energy channels and grades
- Traditional corrections (Ballet, 1999, 2003; Davis, 2001) are deterministic and can fail when...
 - ▶ Pileup is frequent
 - ▶ The source spectrum is complex
 - ▶
- We propose a **fully generative statistical model** for pileup, which enables principled inference

What Actually Happens During Pileup?

- Multiple photons in the same frame \rightarrow Overlapping charge clouds \rightarrow Sum of energies is recorded as a single event with pulse height approximately equal to the sum of energies
- Detector logic applies:
 - ▶ Charge spreading \rightarrow Multiple pixel signal
 - ▶ The center of the pixel island is picked
 - ▶ Grading \rightarrow Determines the measure of spatial distribution of charge
- Observed data: energy channel + event grade (and pixel island location and detector time)

Ingredients

- Within each time frame and detector region (i.e., pixel island) r , we observe...

A channel $C_r \in \{1, \dots, 1024\}$: an observed measure of the photon's (or piled photons') energy with measurement error quantified by a redistribution matrix function

A grade $G_r \in \{\text{good}, \text{bad}\}$ based on the electron charge cloud pattern generated by the photon(s)

- ▶ Events with pileup are usually assigned bad grades due to irregular charge cloud shapes
- ▶ Events with bad grades are typically discarded in practice — but in our model, they are critical for inference!

- We do **not** observe...

P_r , the number of photons *actually* emitted by the source (and then piled up)

$E_{r,1}, \dots, E_{r,P_r}$ the energies of those individual photons

$H_{r,1}, \dots, H_{r,P_r}$ the would-be grades of those individual photons

A Graphical Model

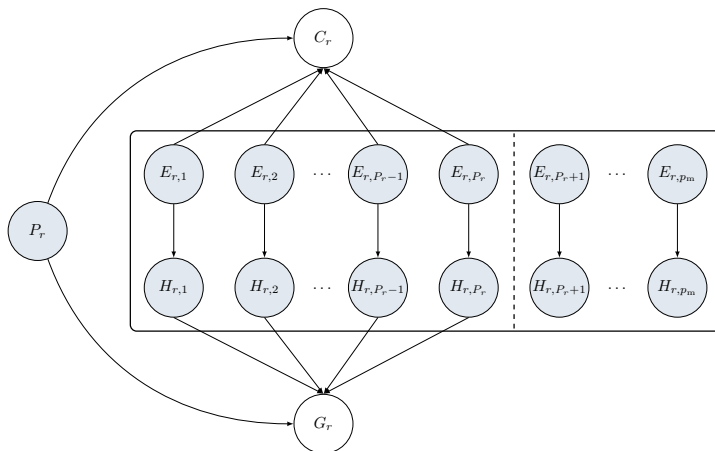


Figure: Graphical model representing the conditional independence structure of the observed and latent variables within a single region r ; shaded nodes represent latent variables

Statistical Inference

- We are mainly interested in the vector of parameters θ corresponding to the source spectrum
- We aim to perform frequentist statistical inference by maximizing the likelihood function

$$L(\theta) = \prod_{r=1}^R \prod_{i=1}^{n_r} \mathbb{P}_{\theta}(G_r = g_{r,i}, C_r = c_{r,i})$$

where $(g_{r,1}, c_{r,1}), \dots, (g_{r,n_r}, c_{r,n_r})$ are independent grade-channel observations in region r

- For this, we need a model that allows us to compute (or estimate) the joint distribution $\mathbb{P}_{\theta}(G_r = g, C_r = c)$

Key Modelling Idea

- Model the number of piled-up photons P_r in region r in one time frame as having a Poisson distribution
- Within a given region r and time frame, let (g, c) be the observed grade-channel pair
- We marginalize over all possible photon configurations and energies:

$$\begin{aligned} & \mathbb{P}_{\boldsymbol{\theta}}(G_r = g, C_r = c) \\ &= \sum_{p=1}^{\infty} \mathbb{P}_{\boldsymbol{\theta}}(G = g, C = c \mid P_r = p) \cdot \mathbb{P}_{\boldsymbol{\theta}}(P_r = p) \\ &= \sum_{p=1}^{\infty} \int \mathbb{P}_{\boldsymbol{\theta}}(G_r = g, C_r = c \mid P_r = p, \mathbf{E}_{1:p} = \mathbf{e}_{1:p}) \cdot q_{p,\boldsymbol{\theta}}(\mathbf{e}_{1:p}) d\mathbf{e}_{1:p} \cdot \mathbb{P}_{\boldsymbol{\theta}}(P_r = p) \end{aligned}$$

- ▶ ...where $q_{p,\boldsymbol{\theta}}(\mathbf{e}_{1:p}) = \prod_{k=1}^p q_{\boldsymbol{\theta}}(e_k)$ is the joint density of the vector of p energies $\mathbf{E}_{1:p} = (E_1, \dots, E_p)$

Complications....

- The distribution of C_r depends on $\mathbf{E}_{1:p}$ through the total energy $\sum_{k=1}^p E_k$, which follows a p -fold convolution of q_θ
- The distribution of the observed grade-channel pair depends on photon-level grades, which are unobserved (so we must marginalize over them too):

$$\begin{aligned} & \mathbb{P}_\theta(G_r = g, C_r = c \mid P_r = p, \mathbf{E}_{1:p} = \mathbf{e}_{1:p}) \\ &= \mathbb{P}_\theta(C_r = c \mid \mathbf{E}_{1:p} = \mathbf{e}_{1:p}) \cdot \sum_{\mathbf{h}_{1:p} \in \{\text{good}, \text{bad}\}^p} \mathbb{P}_\theta(G_r = g \mid \mathbf{H}_{1:p} = \mathbf{h}_{1:p}) \\ & \qquad \qquad \qquad \cdot \mathbb{P}_\theta(\mathbf{H}_{1:p} = \mathbf{h}_{1:p}) \end{aligned}$$

- The detector has a maximum threshold ν_{\max} : if $\sum_{k=1}^p e_k > \nu_{\max}$, then the event is unrecorded (we must account for this too)
- We handle all of these!

Statistical Computation

- The likelihood function can be written down, but it is **analytically intractable** and **impossible to compute exactly** (mainly due to integrals over convolutions)
- Instead, we approximate it by Monte Carlo
- If $E_p^{(1)}, \dots, E_p^{(m)} \stackrel{iid}{\sim} q_{\theta}$ for each $k \geq 1$ and m is large, then

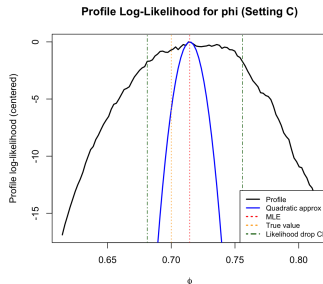
$$\begin{aligned} & \int \mathbb{P}_{\theta}(G_r = g, C_r = c \mid P_r = p, \mathbf{E}_{1:p} = \mathbf{e}_{1:p}) \cdot q_{p,\theta}(\mathbf{e}_{1:p}) d\mathbf{e}_{1:p} \\ & \approx \frac{1}{m} \sum_{j=1}^m \mathbb{P}_{\theta}(G_r = g, C_r = c \mid P_r = p, \mathbf{E}^{(j)}) \end{aligned}$$

- So

$$\begin{aligned} L(\theta) &= \prod_{r=1}^R \prod_{i=1}^{n_r} \mathbb{P}_{\theta}(G_r = g_{r,i}, C_r = c_{r,i}) \\ &\approx \prod_{i=1}^{n_r} \sum_{p=1}^{\infty} \left(\frac{1}{m} \sum_{j=1}^m \mathbb{P}_{\theta}(G_r = g_{r,i}, C_r = c_{r,i} \mid P_r = p, \mathbf{E}^{(j)}) \right) \cdot \mathbb{P}_{\theta}(P_r = p) \end{aligned}$$

Uncertainty Quantification

- In many situations, one can obtain confidence intervals for $\hat{\theta}_{MLE}$ by inspecting the Hessian of $\log(L(\theta))$ at $\hat{\theta}_{MLE}$
- However, our log-likelihood surfaces are locally “wobbly”
 - ▶ We suspect this is due to the discrete nature of observed grades/channels and (maybe) Monte Carlo noise



- Instead, we approximate the log-likelihood surface by dropping points around the MLE and fitting a quadratic surface using linear regression
 - ▶ This (usually) works well!

Simulation Study: Setup

- Write $\lambda(\theta)$ for the total expected photon count per unit frame time
- We simulate data under four scenarios to assess parameter recovery:

Setting A: Low pileup, no emission line

$\lambda(\theta) \approx 2.1$, power law spectrum

Setting B: High pileup, no emission line

$\lambda(\theta) \approx 4.5$, power law spectrum

Setting C: Moderate pileup, mixed spectrum

$\lambda(\theta) \approx 3.3$, power law + emission line spectrum

Setting D: High pileup, mixed spectrum

$\lambda(\theta) \approx 5.6$, power law + emission line spectrum

- These simulations use a fixed RMF matrix and PSF; we fit the model via maximum likelihood using our Monte Carlo likelihood approximation

Simulation Study: Parameter Recovery ($n = 500$)

- Each simulation involves estimating $\theta = (\alpha, \theta)$ or $\theta = (\alpha, \theta, \phi)$, with
 - α : Grade migration parameter which controls probability that piled photons yield a good grade
 - θ : Power law slope
 - ϕ : Relative weight of the power law component vs. the emission line

Setting	α	$\hat{\alpha} \pm \text{SE}$	θ	$\hat{\theta} \pm \text{SE}$	ϕ	$\hat{\phi} \pm \text{SE}$
A	0.70	0.71 ± 0.048	0.70	0.67 ± 0.005	1.00	(fixed)
B	0.70	0.70 ± 0.025	1.50	1.63 ± 0.013	1.00	(fixed)
C	0.70	0.77 ± 0.030	0.70	0.68 ± 0.006	0.70	0.69 ± 0.004
D	0.70	0.68 ± 0.022	1.50	1.63 ± 0.017	0.30	0.28 ± 0.008

- Estimated parameters recover true values well across all conditions
 - Slight upward bias in $\hat{\theta}$ under high pileup (Settings B and D), as expected

Future Directions

- Although we deal with frequentist inference, our work extends easily to the Bayesian regime
 - ▶ Use the same model and use our Monte Carlo based evaluation of the likelihood
- The modular nature of our model allows us to easily add a refined sub-model for the probability of grade migration
- Incorporation of additional instrumental effects
- Accounting for background contamination
- Etc.

Thank you!

References

- Ballet, J. (1999). Pile-up on X-ray CCD instruments. *Astronomy and Astrophysics Supplement Series*, 135(2):371–381.
- Ballet, J. (2003). Pile-up on X-ray CCD instruments. *Advances in Space Research*, 32(10):2077–2082.
- Davis, J. E. (2001). Event Pileup in Charge-coupled Devices. *The Astrophysical Journal*, 562(1):575–582.